Gov 2002 - Causal Inference III: Regression Discontinuity Designs

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- ► IV: instrument provides exogeneous variation
- Regression Discontinuity: exogeneous variation from a discontinuity in treatment assignment

Plan of attack

Sharp Regression Discontinuity Designs

Estimation in the SRD

Readings

Fuzzy Regression Discontinuity Designs

Sharp Regression Discontinuity Designs

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- The classic example of this is in the educational context:
 - Scholarships allocated based on a test score threshold (Thistlethwaite and Campbell, 1960)
 - Class size on test scores using total student thresholds to create new classes (Angrist and Lavy, 1999)

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- Forcing variable: $X_i \in \mathbb{R}$
- Covariates: an *M*-length vector $Z_i = (Z_{1i}, \ldots, Z_{Mi})$

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$$\mathbb{P}(A_i = 1 | X_i = c) = 1$$
$$\mathbb{P}(A_i = 1 | X_i = c - \varepsilon) = 0$$

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$$\lim_{x \downarrow c} E[Y_i | X_i = x] - \lim_{x \uparrow c} E[Y_i | X_i = x]$$

Under certain assumptions, this quantity identifies the ATE at the threshold:

$$\tau_{SRD} = E[Y_i(1) - Y_i(0)|X_i = c]$$

Plotting the RDD (Imbens and Lemieux, 2008)



Fig. 1. Assignment probabilities (SRD).



Fig. 2. Potential and observed outcome regression functions.

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- Thus, we need to extrapolate from the treated to the control group and vice versa.

Extrapolation and smoothness

Remember the quantity of interest here is the effect at the threshold:

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- Extrapolation, even at short distances, requires a certain smoothness in the functions we are extrapolating.

Assumption 1: Continuity

The functions

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Note that this is the same for the treated group:

$$E[Y_i(1)|X_i = c] = \lim_{x \downarrow c} E[Y_i|X_i = x]$$

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Thus, under the ignorability assumption, the sharp RD assumption, and the continuity assumption, we have:

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- ▶ Why? With arbitrarily high N, we'll get an arbitrarily good approximations to the expectation of the line
- How to estimate these nonparametrically is difficult as we'll see (endpoints are a big problem)

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- ► For instance, if people sort around threshold, then you might get jumps other than the one you care about.
- If things other than the treatment change at the threshold, then that might cause discontinuities in the potential outcomes.

Estimation in the SRD

Simple plot of mean outcomes within bins of the forcing variable:

$$\overline{Y}_k = rac{1}{N_k} \sum_{i=1}^N Y_i \cdot \mathbb{I}(b_k < X_i \leq b_{k+1})$$

where N_k is the number of units within bin k and b_k are the bin cutpoints.

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The formal statistical analyses discussed below are essentially just sophisticated versions of this, and if the basic plot does not show any evidence of a discontinuity, there is relatively little chance that the more sophisticated analyses will lead to robust and credible estimates with statistically and substantially significant magnitudes.

Example from RD on extending unemployment



R. Lalive / Journal of Econometrics 142 (2008) 785-806

Discontinuity at threshold = 14.798; with std. err. = 1.928.

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- Similar to balance tests in matching

Checking covariates at the discontinuity



Discontinuity at threshold = 3.442; with std. err. = 1.416.

$$\lim_{x\uparrow c} E[Y_i|X_i=x]$$

The main goal in RD is to estimate the limits of various CEFs such as:

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- Using the entire sample on either side will obviously lead to bias because those values that are far from the cutpoint are clearly different than those nearer to the cutpoint.
- \blacktriangleright \rightarrow restrict our estimation to units close to the threshold.

Example of misleading trends



Let's define

$$\mu_R(x) = \lim_{z \downarrow x} E[Y_i(1)|X_i = z]$$
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- ▶ Here, *h* is a bandwidth parameter, selected by you.
- Basically, calculate means among units no more than h away from the threshold.







• Estimate mean of Y_i when $X_i \in [c, c+h]$ and when $X_i \in [c-h, c)$.

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- Can do this with the following approach regression on those units less than h away from c:

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- This turns out to have very large bias as the we increase the bandwidth.

Local linear regression

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- Instead of a local constant, we can use a local linear regression.
- ► Run a linear regression of Y_i on X_i − c in the group X_i ∈ [c, c + h] to estimate µ₁(x) and the same regression for group with X_i ∈ [c − h, c):

$$(\widehat{\alpha}_L, \widehat{\beta}_L) = \arg\min_{\alpha, \beta} \sum_{i: X_i \in [c-h, c)} (Y_i - \alpha - \beta(X_i - c))^2$$
$$(\widehat{\alpha}_R, \widehat{\beta}_R) = \arg\min_{\alpha, \beta} \sum_{i: X_i \in [c, c+h]} (Y_i - \alpha - \beta(X_i - c))^2$$

Local linear regression

- Instead of a local constant, we can use a local linear regression.
- ► Run a linear regression of Y_i on X_i − c in the group X_i ∈ [c, c + h] to estimate µ₁(x) and the same regression for group with X_i ∈ [c − h, c):

$$(\widehat{\alpha}_L, \widehat{\beta}_L) = \arg\min_{\alpha, \beta} \sum_{i: X_i \in [c-h, c)} (Y_i - \alpha - \beta(X_i - c))^2$$
$$(\widehat{\alpha}_R, \widehat{\beta}_R) = \arg\min_{\alpha, \beta} \sum_{i: X_i \in [c, c+h]} (Y_i - \alpha - \beta(X_i - c))^2$$

Our estimate is

$$egin{aligned} \widehat{ au}_{\mathsf{SRD}} &= \widehat{\mu}_{\mathsf{R}}(\mathsf{c}) - \widehat{\mu}_{\mathsf{L}}(\mathsf{c}) \ &= \widehat{lpha}_{\mathsf{R}} + \widehat{eta}_{\mathsf{R}}(\mathsf{c}-\mathsf{c}) - \widehat{lpha}_{\mathsf{L}} - \widehat{eta}_{\mathsf{L}}(\mathsf{c}-\mathsf{c}) \ &= \widehat{lpha}_{\mathsf{R}} - \widehat{lpha}_{\mathsf{L}} \end{aligned}$$

We can estimate this local linear regression by dropping observations more than h away from c and then running the following regression:

$$Y_i = \alpha + \beta(X_i - c) + \tau A_i + \gamma(X_i - c)A_i + \eta_i$$

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- Here, $\hat{\tau}_{SRD} = \hat{\tau}$ which is the coefficient on the treatment.
- > Yields numerically the same as the separate regressions.

Bandwidth equal to 10 (Global)









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 Make sure that your effects aren't dependent on the polynomial choice.

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- See Imbens and Kalyanaraman (2012) for optimal bandwidth selection.

Readings

Reading 1



Reading 1


Reading 2











Fuzzy Regression Discontinuity Designs

With fuzzy RD, the treatment assignment is no longer a deterministic function of the forcing variable, but there is still a discontinuity in the probability of treatment at the threshold:

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- In the sharp RD, this is also true, but it further required the jump in probability to be from 0 to 1.
- Fuzzy RD is often useful when the a threshold encourages participation in program, but does not actually force units to participate.

Fuzzy RD in a graph





Fig. 4. Potential and observed outcome regression (FRD).



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- $A_i(x) = 1$ if unit *i* would take treatment when X_i was x
- $A_i(x) = 0$ if unit *i* would take control when X_i was x

Assumption 2: Monotoncity

There exists ε such that $A_i(c + e) \ge A_i(c - e)$ for all $0 < e < \varepsilon$

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► Basically, in an *ɛ*-ball around *c*, the forcing variable is randomly assigned.

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- We don't get to see their compliance status because due to the fundamental problem of causal inference
- Could also think about this as changing the threshold instead of changing X_i

Compliance graph



Compliers would not take the treatment if they had X_i = c and we increased the cutoff by some small amount

Compliance graph



- Compliers would not take the treatment if they had X_i = c and we increased the cutoff by some small amount
- These are compliers at the threshold

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LATE in the Fuzzy RD

• We can define an estimator that is in the spirit of IV:

$$\tau_{FRD} = \frac{\lim_{x \downarrow c} E[Y_i | X_i = x] - \lim_{x \uparrow c} E[Y_i | X_i = x]}{\lim_{x \downarrow c} E[A_i | X_i = x] - \lim_{x \uparrow c} E[A_i | X_i = x]}$$
$$= \frac{\text{effect of threshold on } Y_i}{\text{effect of threshold on } A_i}$$

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Under the FRD assumption, continuity, consistency, monotonicity, and local exogeneity, we can write that the estimator is equal to the effect at the threshold for compliers.

$$\tau_{FRD} = \lim_{e \downarrow 0} E[\tau_i | A_i(c+e) > A_i(c-e)]$$

Proof

► To prove this, we'll look at the discontinuity in *Y_i* in a window around the threshold and then shrink that window:

$$E[Y_i|X_i = c + e] - E[Y_i|X_i = c - e]$$
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► Thus, we can write the difference around the threshold as: $E[Y_i|X_i = c+e] - E[Y_i|X_i = c-e] = E[\tau_i(A_i(c+e) - A_i(c-e))]$

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Let's break this expectation apart using the law of iterated expectations:

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$$= E[\tau_i \mid \text{complier}] \times \Pr[\text{complier}]$$

So far, we've shown that the outcome jump at the discontinuity is the LATE times the probability of compliance:

 $E[Y_i|X_i = c + e] - E[Y_i|X_i = c - e] = E[\tau_i \mid \text{complier}] \times \Pr[\text{complier}]$

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Thus,

$$\frac{E[Y_i|X_i = c + e] - E[Y_i|X_i = c - e]}{E[A_i|X_i = c + e] - E[A_i|X_i = c - e]} = E[\tau_i \mid A_i(c+e) > A_i(c-e)]$$

Misc notes

• Taking the limit as $e \rightarrow 0$, we've shown that:

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Note that the FRD estimator emcompasses the SRD estimator because with a sharp design:

$$\lim_{x \downarrow c} E[A_i | X_i = x] - \lim_{x \uparrow c} E[A_i | X_i = x] = 1$$

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Note that the FRD estimator emcompasses the SRD estimator because with a sharp design:

$$\lim_{x \downarrow c} E[A_i | X_i = x] - \lim_{x \uparrow c} E[A_i | X_i = x] = 1$$

A note on external validity: obsviously, FRD puts even more restrictions on the external validity of our estimates because not only are we discussing a LATE, but also the effect is at the threshold. That might give us pause about generalizing other populations for the both the SRD and FRD.

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$$\widehat{\tau}_{FRD} = \frac{\widehat{\tau}_{SRD}}{\widehat{\tau}_{a}}$$

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Thus, being above the threshold is treated like an instrument, controlling for trends in X_i.