Telescope Matching: A Flexible Approach to Estimating Direct Effects

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International Methods Colloquium

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direct effect

effect of treatment not due to a particular downstream cause



effect of treatment not due to a particular downstream cause

why do we care?



effect of treatment not due to a particular downstream cause

why do we care?

causal mediation



effect of treatment not due to a particular downstream cause



direct effect

effect of treatment not due to a particular downstream cause



regression & matching





sequential g-estimation weighting methods





Telecsope matching



consistent for direct effects

Telecsope matching





consistent for direct effectsavoids post-treatment bias





consistent for direct effects
avoids post-treatment bias
robust to (some) model misspecification





consistent for direct effects
avoids post-treatment bias
robust to (some) model misspecification
carries over logic from standard matching

- 1. The difficulty of direct effects
- 2. Our approach: telescope matching
- 3. Simulating misspecification
- 4. Application
- 5. Conclusion

1/ The difficulty of direct effects

Setting

Effect of frame on immigration media accounts

Setting Effect of frame on immigration media accounts A_i Binary treatment ∈ {negative frame, positive frame}







$$\tau(m) = E[Y_i(1, m) - Y_i(0, m)]$$

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 $\tau(m) = E[Y_i(1,m) - Y_i(0,m)]$

- Average effect of manipulating A_i while fixing M_i to level m
- Easily identified if A_i and M_i are randomized but...
- Lots of studies are observational in *M_i* or both.

Confounders



Confounders

baseline covariates



Confounders



Assumptions

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Assumption (Sequential Ignorability)

$$\{Y_i(a, m), M_i(a), Z_i(a)\} \perp A_i | X_i = x$$

$$Y_i(a, m) \perp M_i | A_i = a, X_i = x, Z_i = z$$

No omitted variables for A_i given X_i . No omitted variable for M_i given A_i , X_i , Z_i .

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No omitted variables for
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 given X_i .
No omitted variable for M_i given A_i , X_i , Z_i .

Assumption (Positivity)

$$0 < P(A_i = 1 | X_i = x) < 1$$

$$0 < P(M_i = 1 | X_i = x, Z_i = z, A_i = a) < 1$$

Overlap in the covariate distributions across levels of A_i and M_i

The Problem



The Problem



naive regression/matching of Y_i on X_i , A_i , M_i , and...


naive regression/matching of Y_i on X_i , A_i , M_i , and...

Omit Z_i



naive regression/matching of Y_i on X_i , A_i , M_i , and...





naive regression/matching of Y_i on X_i , A_i , M_i , and...



Control for Z_i



naive regression/matching of Y_i on X_i , A_i , M_i , and...



Extant solutions are model dependent

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Structural Nested Mean Models (SNMMs)

Need the correct model for $\mathbb{E}[Y_i|X_i, A_i, Z_i, M_i]$ and $\mathbb{E}[Y_i|X_i, A_i]$

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Inverse probability of treatment weighting (IPTW)

Need the correct model for $\mathbb{P}[M_i|X_i, A_i, Z_i]$ and $\mathbb{P}[A_i|X_i]$

2/ Our approach: telescope matching







Two-stage matching procedure

Match M_i on Z_i , A_i , and X_i





Telescope matching



Telescope matching



Unit		Obse	erved		Potential Outcomes				
	A_i	M_i	X_i	Z_i	$Y_i(1, 1)$	$Y_i(1, 0)$	$Y_i(0, 1)$	$Y_i(0, 0)$	

Unit	Observed				Potential Outcomes				
	A_i	M_i	X_i	Z_i	$Y_i(1, 1)$	$Y_i(1, 0)$	$Y_i(0, 1)$	$Y_i(0, 0)$	
1	1	1	10	3	Y ₁	?	?	?	

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1	1	1	10	3	Y ₁	?	?	?
2	1	0	9	2	?	Y_2	?	?

Unit		Obse	erved		Potential Outcomes			
	A_i	M_i	X_i	Z_i	$Y_i(1, 1)$	$Y_i(1, 0)$	<i>Y_i</i> (0, 1)	$Y_i(0, 0)$
1	1	1	10	3	<i>Y</i> ₁	?	?	?
2	1	0	9	2	?	Y_2	?	?
3	1	0	8	1	?	Y_3	?	?
4	0	1	8	3	?	?	Y_4	?
5	0	0	9	2	?	?	?	Y_5
6	0	0	10	1	?	?	?	Y ₆

Unit		Observed Potential Outcomes					es	
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1	1	1	10	3	Y_1	?	?	?
2	1	0	9	2	?	Y_2	?	?
3	1	0	8	1	?	Y_3	?	?
4	0	1	8	3	?	?	Y_4	?
5	0	0	9	2	?	?	?	Y_5
6	0	0	10	1	?	?	?	Y ₆

 $\tau(0) = E[Y_i(1,0) - Y_i(0,0)]$









1. Subset to a particular level of A_i 2. Match each $M_i = 1$ to closest $M_i = 0$ unit in $\{X_{ii}, Z_i\}$





3. Impute missing counterfactual with matched \hat{Y}_{i0}

$$A_{i} = 0 \qquad A_{i} = 0
M_{i} = 1 \qquad M_{i} = 0
Y_{1} \qquad Y_{2} \rightarrow \widehat{Y}_{1,0} = Y_{2} \approx Y_{1}(0,0)
Y_{3} \qquad Y_{4} \rightarrow \widehat{Y}_{3,0} = Y_{4} \approx Y_{3}(0,0)
Y_{5} \qquad Y_{6} \rightarrow \widehat{Y}_{5,0} = Y_{6} \approx Y_{5}(0,0)$$



1:1 matching example

Unit		Obse	erved		Potential Outcomes				
	A_i	M_i	X_i	Z_i	$Y_i(1, 1)$	<i>Y_i</i> (1, 0)	$Y_i(0, 1)$	$Y_i(0, 0)$	
1	1	1	10	3	Y_1	?	?	?	
2	1	0	9	2	?	Y_2	?	?	
3	1	0	8	1	?	Y_3	?	?	
4	0	1	8	3	?	?	Y_4	?	
5	0	0	9	2	?	?	?	Y_5	
6	0	0	10	1	?	?	?	Y ₆	

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1	1	1	10	3	Y_1	?	?	?	
2	1	0	9	2	?	<i>Y</i> ₂	?	?	
3	1	0	8	1	?	Y ₃	?	?	
4	0	1	8	3	?	?	Y_4	?	
5	0	0	9	2	?	?	?	Y_5	
6	0	0	10	1	?	?	?	Y_6	

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1	1	1	10	3	Y_1	Y_2	?	?
2	1	0	9	2	?	Y_2	?	?
3	1	0	8	1	?	Y_3	?	?
4	0	1	8	3	?	?	Y_4	?
5	0	0	9	2	?	?	?	Y_5
6	0	0	10	1	?	?	?	Y ₆

Standard matching using \hat{Y}_{i0} as outcome completely ignoring M_i and Z_i

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Standard matching using \hat{Y}_{i0} as outcome completely ignoring M_i and Z_i

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Standard matching using \hat{Y}_{i0} as outcome completely ignoring M_i and Z_i

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1:1 matching, second stage

Unit	Observed				Potential Outcomes			
	A_i	M_i	X_i	Z_i	$Y_i(1, 1)$	$Y_i(1, 0)$	$Y_i(0, 1)$	$Y_i(0, 0)$
1	1	1	10	3	Y_1	Y_2	?	?
2	1	0	9	2	?	Y_2	?	?
3	1	0	8	1	?	Y ₃	?	?
4	0	1	8	3	?	?	Y_4	?
5	0	0	9	2	?	?	?	Y_5
6	0	0	10	1	?	?	?	Y ₆

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1	1	1	10	3	Y_1	Y_2	?	?
2	1	0	9	2	?	Y_2	?	?
3	1	0	8	1	?	Y_3	?	?
4	0	1	8	3	?	Y ₃	Y_4	?
5	0	0	9	2	?	?	?	Y_5
6	0	0	10	1	?	?	?	Y ₆
Unit	Observed				Potential Outcomes			
------	----------	-------	-------	-------	--------------------	-----------------------------	-------------	----------------
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1	1	1	10	3	Y_1	Y_2	?	?
2	1	0	9	2	?	$Y_2 \prec$?	?
3	1	0	8	1	?	Y_3	?	?
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2	1	0	9	2	?	Y_2	?	Y_5
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6	0	0	10	1	?	$\searrow Y_2$?	Y ₆

$$\widehat{\tau} = \frac{1}{6} \Big[(Y_2 - Y_6) + (Y_2 - Y_5) + (Y_3 - Y_5) \\ + (Y_3 - Y_5) + (Y_2 - Y_5) + (Y_2 - Y_6) \Big]$$

Simple



two standard matching steps

Simple

•two standard matching steps
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velop bias correction





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Weighted bootstrap

•we follow Otsu and Rai (2017) and resample each contribution to the estimator

3/ Simulating misspecification





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- Effect of A only through M so true ACDE: $\tau(0) = 0$
- δ controls magnitude of post-treatment confounding
 - ▶ when $\delta \neq 0$, controlling for Z in a naive regression will induce post-treatment bias.

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Number of matches per stage = 3







Simulation results: Root Mean Square Error



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4/ Application

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- Pre-treatment confounders (X_i) : Education, Gender, Income, Age.
- Post-treatment confounder (Z_i) : Perceived harm due to immigration.







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- Telescope matching shows ACDE closer to zero, high uncertainty.
- \rightsquigarrow Fixing the mediator eliminates most of the treatment effect.

5/ Conclusion

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- Direct effects models such as SNMMs and IPTW are model dependent.
- We introduce two-stage matching procedure to close this gap.
 - Estimator is consistent, but biased, so we use bias correction.
 - Weighted bootstrap for uncertainty estimates.

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- Apply ideas to mediation analysis where there are no Z_i .

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- Apply ideas to mediation analysis where there are no Z_i.
- How to handle dropping units in the first stage since it induces post-treatment bias?

- For more information, see:
- •

http://www.mattblackwell.org/files/papers/telescope_matching

- http://www.mattblackwell.org
- https://www.antonstrezhnev.com/

SNMMs as imputation estimators

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1. Estimate the conditional effect of M_i
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$$\overline{Y}_i(A_i, 0) = Y_i - M_i \times \widehat{\gamma}_m(X_i, Z_i, A_i)$$

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2. Impute $Y_i(A_i, 0)$ by subtracting effect of M_i

$$\overline{Y}_i(A_i, 0) = Y_i - M_i \times \widehat{\gamma}_m(X_i, Z_i, A_i)$$

3. Regress imputations on treatment and baseline covariates to get ACDE

1. Estimate the conditional effect of M_i

$$\gamma_m(x, z, a) = \mathbb{E}[Y_i(a, 1) - Y_i(a, 0) | X_i = x, A_i = a, Z_i = z, M_i = 1]$$

2. Impute $Y_i(A_i, 0)$ by subtracting effect of M_i

$$\overline{Y}_i(A_i, 0) = Y_i - M_i \times \widehat{\gamma}_m(X_i, Z_i, A_i)$$

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 $\mathbb{E}[\overline{Y}_i(A_i, 0)|X_i, A_i]$

Depends on correct model for $E[Y_i|X_i, Z_i, A_i, M_i]$ and $E[Y_i|X_i, A_i]$