

# Gov 2000: 9. Multiple Regression in Matrix Form

Matthew Blackwell

November 19, 2015

1. Matrix algebra review
2. Matrix Operations
3. Writing the linear model more compactly
4. A bit more about matrices
5. OLS in matrix form
6. OLS inference in matrix form

# Where are we? Where are we going?

- Last few weeks: regression estimation and inference with one and two independent variables, varying effects
- This week: the general regression model with arbitrary covariates
- Next week: what happens when assumptions are wrong

# Nunn & Wantchekon

- Are there long-term, persistent effects of slave trade on Africans today?
- Basic idea: compare levels of interpersonal trust ( $Y_i$ ) across different levels of historical slave exports for a respondent's ethnic group
- Problem: ethnic groups and respondents might differ in their interpersonal trust in ways that correlate with the severity of slave exports
- One solution: try to control for relevant differences between groups via multiple regression

### III. Estimating Equations and Empirical Results

#### A. OLS Estimates

We begin by estimating the relationship between the number of slaves that were taken from an individual's ethnic group and the individual's current level of trust. Our baseline estimating equation is:

$$(1) \text{ trust}_{i,e,d,c} = \alpha_c + \beta \text{ slave exports}_e + \mathbf{X}'_{i,e,d,c} \mathbf{\Gamma} + \mathbf{X}'_{d,c} \mathbf{\Omega} + \mathbf{X}'_e \mathbf{\Phi} + \varepsilon_{i,e,d,c}$$

- Whaaaaa? Bold letter, quotation marks, what is this?
- Today's goal is to decipher this type of writing

# Multiple Regression in R

```
nunn <- foreign::read.dta("Nunn_Wantchekon_AER_2011.dta")
mod <- lm(trust_neighbors ~ exports + age + male + urban_dum
          + malaria_ecology, data = nunn)
summary(mod)
```

```
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   1.5030370  0.0218325   68.84  <2e-16 ***
## exports      -0.0010208  0.0000409  -24.94  <2e-16 ***
## age           0.0050447  0.0004724   10.68  <2e-16 ***
## male         0.0278369  0.0138163    2.01   0.044 *
## urban_dum    -0.2738719  0.0143549  -19.08  <2e-16 ***
## malaria_ecology 0.0194106  0.0008712   22.28  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.978 on 20319 degrees of freedom
## (1497 observations deleted due to missingness)
## Multiple R-squared:  0.0604, Adjusted R-squared:  0.0602
## F-statistic: 261 on 5 and 20319 DF, p-value: <2e-16
```

# Why matrices and vectors?





# Why matrices and vectors?

- Here's one way to write the full multiple regression model:

$$y_i = \beta_0 + x_{i1}\beta_1 + x_{i2}\beta_2 + \cdots + x_{ik}\beta_k + u_i$$

- Notation is going to get needlessly messy as we add variables
- Matrices are clean, but they are like a foreign language
- You need to build intuitions over a long period of time

# Quick note about interpretation

$$y_i = \beta_0 + x_{i1}\beta_1 + x_{i2}\beta_2 + \cdots + x_{ik}\beta_k + u_i$$

- In this model,  $\beta_1$  is the effect of a one-unit change in  $x_{i1}$  conditional on all other  $x_{ij}$ .
- Jargon “partial effect,” “ceteris paribus,” “all else equal,” “conditional on the covariates,” etc

# 1/ Matrix algebra review

# Matrices and vectors

- A matrix is just a rectangular array of numbers.
- We say that a matrix is  $n \times k$  (“ $n$  by  $k$ ”) if it has  $n$  rows and  $k$  columns.
- Uppercase bold denotes a matrix:

$$\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} & \cdots & a_{1k} \\ a_{21} & a_{22} & \cdots & a_{2k} \\ \vdots & \vdots & \ddots & \vdots \\ a_{n1} & a_{n2} & \cdots & a_{nk} \end{bmatrix}$$

- Generic entry:  $a_{ij}$  where this is the entry in row  $i$  and column  $j$
- If you've used Excel, you've seen matrices.

# Examples of matrices

- One example of a matrix that we'll use a lot is the **design matrix**, which has a column of ones, and then each of the subsequent columns is each independent variable in the regression.

$$\mathbf{X} = \begin{bmatrix} 1 & \text{exports}_1 & \text{age}_1 & \text{male}_1 \\ 1 & \text{exports}_2 & \text{age}_2 & \text{male}_2 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & \text{exports}_n & \text{age}_n & \text{male}_n \end{bmatrix}$$

# Design matrix in R

```
head(model.matrix(mod), 8)
```

```
##   (Intercept) exports age male urban_dum malaria_ecology
## 1           1     855  40   0         0           28.15
## 2           1     855  25   1         0           28.15
## 3           1     855  38   1         1           28.15
## 4           1     855  37   0         1           28.15
## 5           1     855  31   1         0           28.15
## 6           1     855  45   0         0           28.15
## 7           1     855  20   1         0           28.15
## 8           1     855  31   0         0           28.15
```

```
dim(model.matrix(mod))
```

```
## [1] 20325    6
```

# Vectors

- A **vector** is just a matrix with only one row or one column.
- A **row vector** is a vector with only one row, sometimes called a  $1 \times k$  vector:

$$\boldsymbol{\alpha} = [ \alpha_1 \quad \alpha_2 \quad \alpha_3 \quad \cdots \quad \alpha_k ]$$

- A **column vector** is a vector with one column and more than one row. Here is a  $n \times 1$  vector:

$$\mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$

- **Convention** we'll assume that a vector is column vector and vectors will be written with lowercase bold lettering (**b**)

# Vector examples

- One really common vector that we will work with are individual variables, such as the dependent variable, which we will represent as  $\mathbf{y}$ :

$$\mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$



# Vectors in R

- Vectors can come from subsets of matrices:

```
model.matrix(mod)[1, ]
```

```
##      (Intercept)      exports      age
##           1.00      854.96    40.00
##      urban_dum malaria_ecology
##           0.00      28.15
```

- Note, though, that R always prints a vector in row form, even if it is a column in the original data:

```
head(nunn$trust_neighbors)
```

```
## [1] 3 3 0 0 1 1
```

# 2/ Matrix Operations

# Transpose

- The **transpose** of a matrix  $\mathbf{A}$  is the matrix created by switching the rows and columns of the data and is denoted  $\mathbf{A}'$ .
- $k$ th column of  $\mathbf{A}$  becomes the  $k$ th row of  $\mathbf{A}'$ :

$$\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \\ a_{31} & a_{32} \end{bmatrix} \quad \mathbf{A}' = \begin{bmatrix} a_{11} & a_{21} & a_{31} \\ a_{12} & a_{22} & a_{32} \end{bmatrix}$$

- If  $\mathbf{A}$  is  $n \times k$ , then  $\mathbf{A}'$  will be  $k \times n$ .
- Also written  $\mathbf{A}^T$

# Transposing vectors

- Transposing will turn a  $k \times 1$  column vector into a  $1 \times k$  row vector and vice versa:

$$\omega = \begin{bmatrix} 1 \\ 3 \\ 2 \\ -5 \end{bmatrix} \quad \omega' = [ 1 \quad 3 \quad 2 \quad -5 ]$$

# Transposing in R

```
a <- matrix(1:6, ncol = 3, nrow = 2)
a
```

```
##      [,1] [,2] [,3]
## [1,]    1    3    5
## [2,]    2    4    6
```

```
t(a)
```

```
##      [,1] [,2]
## [1,]    1    2
## [2,]    3    4
## [3,]    5    6
```

# Write matrices as vectors

- A matrix is just a collection of vectors (row or column)
- As a row vector:

$$\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \end{bmatrix} = \begin{bmatrix} \mathbf{a}'_1 \\ \mathbf{a}'_2 \end{bmatrix}$$

with row vectors

$$\mathbf{a}'_1 = [ a_{11} \quad a_{12} \quad a_{13} ] \quad \mathbf{a}'_2 = [ a_{21} \quad a_{22} \quad a_{23} ]$$

- Or we can define it in terms of column vectors:

$$\mathbf{B} = \begin{bmatrix} b_{11} & b_{12} \\ b_{21} & b_{22} \\ b_{31} & b_{32} \end{bmatrix} = [ \mathbf{b}_1 \quad \mathbf{b}_2 ]$$

where  $\mathbf{b}_1$  and  $\mathbf{b}_2$  represent the columns of  $\mathbf{B}$ .

- $j$  subscripts columns of a matrix:  $\mathbf{x}_j$
- $i$  and  $t$  will be used for rows  $\mathbf{x}'_i$ .

# Addition and subtraction

- How do we add or subtract matrices and vectors?
- First, the matrices/vectors need to be **comformable**, meaning that the dimensions have to be the same
- Let **A** and **B** both be  $2 \times 2$  matrices. Then, let **C** = **A** + **B**, where we add each cell together:

$$\begin{aligned}\mathbf{A} + \mathbf{B} &= \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix} + \begin{bmatrix} b_{11} & b_{12} \\ b_{21} & b_{22} \end{bmatrix} \\ &= \begin{bmatrix} a_{11} + b_{11} & a_{12} + b_{12} \\ a_{21} + b_{21} & a_{22} + b_{22} \end{bmatrix} \\ &= \begin{bmatrix} c_{11} & c_{12} \\ c_{21} & c_{22} \end{bmatrix} \\ &= \mathbf{C}\end{aligned}$$

# Scalar multiplication

- A scalar is just a single number: you can think of it sort of like a 1 by 1 matrix.
- When we multiply a scalar by a matrix, we just multiply each element/cell by that scalar:

$$\alpha \mathbf{A} = \alpha \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix} = \begin{bmatrix} \alpha \times a_{11} & \alpha \times a_{12} \\ \alpha \times a_{21} & \alpha \times a_{22} \end{bmatrix}$$



# **3/** Writing the linear model more compactly

# The linear model with new notation

- Remember that we wrote the linear model as the following for all  $i \in [1, \dots, n]$ :

$$y_i = \beta_0 + x_i\beta_1 + z_i\beta_2 + u_i$$

- Imagine we had an  $n$  of 4. We could write out each formula:

$$y_1 = \beta_0 + x_1\beta_1 + z_1\beta_2 + u_1 \quad (\text{unit 1})$$

$$y_2 = \beta_0 + x_2\beta_1 + z_2\beta_2 + u_2 \quad (\text{unit 2})$$

$$y_3 = \beta_0 + x_3\beta_1 + z_3\beta_2 + u_3 \quad (\text{unit 3})$$

$$y_4 = \beta_0 + x_4\beta_1 + z_4\beta_2 + u_4 \quad (\text{unit 4})$$

# The linear model with new notation

$$y_1 = \beta_0 + x_1\beta_1 + z_1\beta_2 + u_1 \quad (\text{unit 1})$$

$$y_2 = \beta_0 + x_2\beta_1 + z_2\beta_2 + u_2 \quad (\text{unit 2})$$

$$y_3 = \beta_0 + x_3\beta_1 + z_3\beta_2 + u_3 \quad (\text{unit 3})$$

$$y_4 = \beta_0 + x_4\beta_1 + z_4\beta_2 + u_4 \quad (\text{unit 4})$$

- We can write this as:

$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{bmatrix} = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} \beta_0 + \begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \end{bmatrix} \beta_1 + \begin{bmatrix} z_1 \\ z_2 \\ z_3 \\ z_4 \end{bmatrix} \beta_2 + \begin{bmatrix} u_1 \\ u_2 \\ u_3 \\ u_4 \end{bmatrix}$$

- Outcome is a linear combination of the the  $\mathbf{x}$ ,  $\mathbf{z}$ , and  $\mathbf{u}$  vectors

# Grouping things into matrices

- Can we write this in a more compact form? Yes! Let  $\mathbf{X}$  and  $\boldsymbol{\beta}$  be the following:

$$\mathbf{X}_{(4 \times 3)} = \begin{bmatrix} 1 & x_1 & z_1 \\ 1 & x_2 & z_2 \\ 1 & x_3 & z_3 \\ 1 & x_4 & z_4 \end{bmatrix} \quad \boldsymbol{\beta}_{(3 \times 1)} = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{bmatrix}$$

# Matrix multiplication by a vector

- We can write this more compactly as a matrix (post-)multiplied by a vector:

$$\begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} \beta_0 + \begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \end{bmatrix} \beta_1 + \begin{bmatrix} z_1 \\ z_2 \\ z_3 \\ z_4 \end{bmatrix} \beta_2 = \mathbf{X}\boldsymbol{\beta}$$

- Multiplication of a matrix by a vector is just the **linear combination** of the columns of the matrix with the vector elements as weights/coefficients.
- And the left-hand side here only uses scalars times vectors, which is easy!

# General matrix by vector multiplication

- $\mathbf{A}$  is a  $n \times k$  matrix
- $\mathbf{b}$  is a  $k \times 1$  column vector
- Columns of  $\mathbf{A}$  have to match rows of  $\mathbf{b}$
- Let  $\mathbf{a}_j$  be the  $j$ th column of  $A$ . Then we can write:

$$\underset{(n \times 1)}{\mathbf{c}} = \mathbf{A}\mathbf{b} = b_1\mathbf{a}_1 + b_2\mathbf{a}_2 + \cdots + b_k\mathbf{a}_k$$

- $\mathbf{c}$  is linear combination of the columns of  $\mathbf{A}$

# Back to regression

- $\mathbf{X}$  is the  $n \times (k + 1)$  design matrix of independent variables
- $\boldsymbol{\beta}$  be the  $(k + 1) \times 1$  column vector of coefficients.
- $\mathbf{X}\boldsymbol{\beta}$  will be  $n \times 1$ :

$$\mathbf{X}\boldsymbol{\beta} = \beta_0 + \beta_1\mathbf{x}_1 + \beta_2\mathbf{x}_2 + \cdots + \beta_k\mathbf{x}_k$$

- Thus, we can compactly write the linear model as the following:

$$\underset{(n \times 1)}{\mathbf{y}} = \underset{(n \times 1)}{\mathbf{X}\boldsymbol{\beta}} + \underset{(n \times 1)}{\mathbf{u}}$$

- We can also write this at the individual level, where  $\mathbf{x}'_i$  is the  $i$ th row of  $\mathbf{X}$ :

$$y_i = \mathbf{x}'_i\boldsymbol{\beta} + u_i$$

# 4/ A bit more about matrices



# Matrix multiplication

- What if, instead of a column vector  $b$ , we have a matrix  $\mathbf{B}$  with dimensions  $k \times m$ .
- How do we do multiplication like so  $\mathbf{C} = \mathbf{AB}$ ?
- Each column of the new matrix is just matrix by vector multiplication:

$$\mathbf{C} = [\mathbf{c}_1 \quad \mathbf{c}_2 \quad \cdots \quad \mathbf{c}_m] \quad \mathbf{c}_j = \mathbf{A}\mathbf{b}_j$$

- Thus, each column of  $\mathbf{C}$  is a linear combination of the columns of  $\mathbf{A}$ .

# Special multiplications

- The **inner product** of a two column vectors **a** and **b** (of equal dimension,  $k \times 1$ ):

$$\mathbf{a}'\mathbf{b} = a_1b_1 + a_2b_2 + \cdots + a_kb_k$$

- Special case of above: **a'** is a matrix with  $k$  columns and just 1 row, so the “columns” of **a'** are just scalars.

# Sum of the squared residuals

- Example: let's say that we have a vector of residuals,  $\hat{\mathbf{u}}$ , then the inner product of the residuals is:

$$\hat{\mathbf{u}}' \hat{\mathbf{u}} = \begin{bmatrix} \hat{u}_1 & \hat{u}_2 & \cdots & \hat{u}_n \end{bmatrix} \begin{bmatrix} \hat{u}_1 \\ \hat{u}_2 \\ \vdots \\ \hat{u}_n \end{bmatrix}$$

$$\hat{\mathbf{u}}' \hat{\mathbf{u}} = \hat{u}_1 \hat{u}_1 + \hat{u}_2 \hat{u}_2 + \cdots + \hat{u}_n \hat{u}_n = \sum_{i=1}^n \hat{u}_i^2$$

- It's just the sum of the squared residuals!

# Square matrices and the diagonal

- A **square matrix** has equal numbers of rows and columns.
- The **diagonal** of a square matrix are the values  $a_{jj}$ :

$$\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \\ a_{31} & a_{32} & a_{33} \end{bmatrix}$$

- The **identity matrix**,  $\mathbf{I}$  is a square matrix, with 1s along the diagonal and 0s everywhere else.

$$\mathbf{I} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

- The identity matrix multiplied by any matrix returns the matrix:  $\mathbf{AI} = \mathbf{A}$ .

# Identity matrix

- To get the diagonal of a matrix in R, use the `diag()` function:

```
b <- matrix(1:4, nrow = 2, ncol = 2)
b
```

```
##      [,1] [,2]
## [1,]    1    3
## [2,]    2    4
```

```
diag(b)
```

```
## [1] 1 4
```

- `diag()` also creates identity matrices in R:

```
diag(3)
```

```
##      [,1] [,2] [,3]
## [1,]    1    0    0
## [2,]    0    1    0
## [3,]    0    0    1
```

# 5/ OLS in matrix form

# Multiple linear regression in matrix form

- Let  $\widehat{\boldsymbol{\beta}}$  be the matrix of estimated regression coefficients and  $\widehat{\mathbf{y}}$  be the vector of fitted values:

$$\widehat{\boldsymbol{\beta}} = \begin{bmatrix} \widehat{\beta}_0 \\ \widehat{\beta}_1 \\ \vdots \\ \widehat{\beta}_k \end{bmatrix} \quad \widehat{\mathbf{y}} = \mathbf{X}\widehat{\boldsymbol{\beta}}$$

- It might be helpful to see this again more written out:

$$\widehat{\mathbf{y}} = \begin{bmatrix} \widehat{y}_1 \\ \widehat{y}_2 \\ \vdots \\ \widehat{y}_n \end{bmatrix} = \mathbf{X}\widehat{\boldsymbol{\beta}} = \begin{bmatrix} 1\widehat{\beta}_0 + x_{11}\widehat{\beta}_1 + x_{12}\widehat{\beta}_2 + \cdots + x_{1k}\widehat{\beta}_k \\ 1\widehat{\beta}_0 + x_{21}\widehat{\beta}_1 + x_{22}\widehat{\beta}_2 + \cdots + x_{2k}\widehat{\beta}_k \\ \vdots \\ 1\widehat{\beta}_0 + x_{n1}\widehat{\beta}_1 + x_{n2}\widehat{\beta}_2 + \cdots + x_{nk}\widehat{\beta}_k \end{bmatrix}$$

# Residuals

- We can easily write the residuals in matrix form:

$$\hat{\mathbf{u}} = \mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}$$

- Our goal as usual is to minimize the sum of the squared residuals, which we saw earlier we can write:

$$\hat{\mathbf{u}}'\hat{\mathbf{u}} = (\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}})'(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}})$$



# OLS estimator in matrix form

- Goal: minimize the sum of the squared residuals
- Take (matrix) derivatives, set equal to 0 (see Wooldridge for details)
- Resulting first order conditions:

$$\mathbf{X}'(\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}) = 0$$

- Rearranging:

$$\mathbf{X}'\mathbf{X}\hat{\boldsymbol{\beta}} = \mathbf{X}'\mathbf{y}$$

- In order to isolate  $\hat{\boldsymbol{\beta}}$ , we need to move the  $\mathbf{X}'\mathbf{X}$  term to the other side of the equals sign.
- We've learned about matrix multiplication, but what about matrix "division"?

# Scalar inverses

- What is division in its simplest form?  $\frac{1}{a}$  is the value such that  $a\frac{1}{a} = 1$ :
- For some algebraic expression:  $au = b$ , let's solve for  $u$ :

$$\frac{1}{a}au = \frac{1}{a}b$$
$$u = \frac{b}{a}$$

- Need a matrix version of this:  $\frac{1}{a}$ .

# Matrix inverses

- **Definition** If it exists, the **inverse** of square matrix  $\mathbf{A}$ , denoted  $\mathbf{A}^{-1}$ , is the matrix such that  $\mathbf{A}^{-1}\mathbf{A} = \mathbf{I}$ .
- We can use the inverse to solve (systems of) equations:

$$\mathbf{A}\mathbf{u} = \mathbf{b}$$

$$\mathbf{A}^{-1}\mathbf{A}\mathbf{u} = \mathbf{A}^{-1}\mathbf{b}$$

$$\mathbf{I}\mathbf{u} = \mathbf{A}^{-1}\mathbf{b}$$

$$\mathbf{u} = \mathbf{A}^{-1}\mathbf{b}$$

- If the inverse exists, we say that  $\mathbf{A}$  is **invertible** or **nonsingular**.

# Back to OLS

- Let's assume, for now, that the inverse of  $\mathbf{X}'\mathbf{X}$  exists (we'll come back to this)
- Then we can write the OLS estimator as the following:

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$$

- Memorize this: “ex prime ex inverse ex prime y” sear it into your soul.

# OLS by hand in R

$$\widehat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$$

- First we need to get the design matrix and the response:

```
X <- model.matrix(trust_neighbors ~ exports + age + male
                  + urban_dum + malaria_ecology, data = nunn)
dim(X)
```

```
## [1] 20325      6
```

```
## model.frame always puts the response in the first column
y <- model.frame(trust_neighbors ~ exports + age + male
                 + urban_dum + malaria_ecology, data = nunn)[,1]
length(y)
```

```
## [1] 20325
```

# OLS by hand in R

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$$

- Use the `solve()` for inverses and `%*%` for matrix multiplication:

```
solve(t(X) %*% X) %*% t(X) %*% y
```

```
##      (Intercept)  exports      age      male urban_dum
## [1,]      1.503 -0.001021  0.005045  0.02784   -0.2739
##      malaria_ecology
## [1,]      0.01941
```

```
coef(mod)
```

```
##      (Intercept)      exports      age      male
##      1.503037      -0.001021      0.005045      0.027837
##      urban_dum malaria_ecology
##      -0.273872      0.019411
```

# Intuition for the OLS in matrix form

$$\widehat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$$

- What's the intuition here?
- “Numerator”  $\mathbf{X}'\mathbf{y}$ : is roughly composed of the covariances between the columns of  $\mathbf{X}$  and  $\mathbf{y}$
- “Denominator”  $\mathbf{X}'\mathbf{X}$  is roughly composed of the sample variances and covariances of variables within  $\mathbf{X}$
- Thus, we have something like:

$$\widehat{\boldsymbol{\beta}} \approx (\text{variance of } \mathbf{X})^{-1}(\text{covariance of } \mathbf{X} \text{ \& } \mathbf{y})$$

- This is a rough sketch and isn't strictly true, but it can provide intuition.

# **6/** OLS inference in matrix form



# Random vectors

- A **random vector** is a vector of random variables:

$$\mathbf{x}_i = \begin{bmatrix} x_{i1} \\ x_{i2} \end{bmatrix}$$

- Here,  $\mathbf{x}_i$  is a random vector and  $x_{i1}$  and  $x_{i2}$  are random variables.
- When we talk about the distribution of  $\mathbf{x}_i$ , we are talking about the joint distribution of  $x_{i1}$  and  $x_{i2}$ .

# Distribution of random vectors

$$\mathbf{x}_i \sim (\mathbb{E}[\mathbf{x}_i], \mathbb{V}[\mathbf{x}_i])$$

- Expectation of random vectors:

$$\mathbb{E}[\mathbf{x}_i] = \begin{bmatrix} \mathbb{E}[x_{i1}] \\ \mathbb{E}[x_{i2}] \end{bmatrix}$$

- Variance of random vectors:

$$\mathbb{V}[\mathbf{x}_i] = \begin{bmatrix} \mathbb{V}[x_{i1}] & \text{Cov}[x_{i1}, x_{i2}] \\ \text{Cov}[x_{i1}, x_{i2}] & \mathbb{V}[x_{i2}] \end{bmatrix}$$

- Variance of the random vector also called the [variance-covariance matrix](#).
- These describe the joint distribution of  $\mathbf{x}_i$

# Most general OLS assumptions

1. Linearity:  $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{u}$
2. Random/iid sample:  $(y_i, \mathbf{x}'_i)$  are a iid sample from the population.
3. No perfect collinearity:  $\mathbf{X}$  is an  $n \times (k + 1)$  matrix with rank  $k + 1$
4. Zero conditional mean:  $\mathbb{E}[\mathbf{u}|\mathbf{X}] = \mathbf{0}$
5. Homoskedasticity:  $\text{var}(\mathbf{u}|\mathbf{X}) = \sigma_u^2 \mathbf{I}_n$
6. Normality:  $\mathbf{u}|\mathbf{X} \sim N(\mathbf{0}, \sigma_u^2 \mathbf{I}_n)$

# No perfect collinearity

- In matrix form:  $\mathbf{X}$  is an  $n \times (k + 1)$  matrix with rank  $k + 1$
- **Definition** The **rank** of a matrix is the maximum number of linearly independent columns.
- If  $\mathbf{X}$  has rank  $k + 1$ , then all of its columns are linearly independent
- ...and none of its columns are linearly dependent  $\implies$  no perfect collinearity
- $\mathbf{X}$  has rank  $k + 1 \rightsquigarrow (\mathbf{X}'\mathbf{X})$  is invertible
- Just like variation in  $X_i$  led us to be able to divide by the variance in simple OLS

# Zero conditional mean error

- Using the zero mean conditional error assumptions:

$$\mathbb{E}[\mathbf{u}|\mathbf{X}] = \begin{bmatrix} \mathbb{E}[u_1|\mathbf{X}] \\ \mathbb{E}[u_2|\mathbf{X}] \\ \vdots \\ \mathbb{E}[u_n|\mathbf{X}] \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} = \mathbf{0}$$

# OLS is unbiased

- Under matrix assumptions 1-4, OLS is unbiased for  $\beta$ :

$$\mathbb{E}[\widehat{\beta}] = \beta$$

# Variance-covariance matrix

- The homoskedasticity assumption is different:

$$\text{var}(\mathbf{u}|\mathbf{X}) = \sigma_u^2 \mathbf{I}_n$$

- In order to investigate this, we need to know what the variance of a vector is.
- The variance of a vector is actually a matrix:

$$\text{var}[\mathbf{u}] = \Sigma_u = \begin{bmatrix} \text{var}(u_1) & \text{cov}(u_1, u_2) & \dots & \text{cov}(u_1, u_n) \\ \text{cov}(u_2, u_1) & \text{var}(u_2) & \dots & \text{cov}(u_2, u_n) \\ \vdots & & \ddots & \\ \text{cov}(u_n, u_1) & \text{cov}(u_n, u_2) & \dots & \text{var}(u_n) \end{bmatrix}$$

- This matrix is symmetric since  $\text{cov}(u_i, u_j) = \text{cov}(u_j, u_i)$

# Matrix version of homoskedasticity

- Once again:  $\text{var}(\mathbf{u}|\mathbf{X}) = \sigma_u^2 \mathbf{I}_n$
- $\mathbf{I}_n$  is the  $n \times n$  identity matrix
- Visually:

$$\text{var}[\mathbf{u}] = \sigma_u^2 \mathbf{I}_n = \begin{bmatrix} \sigma_u^2 & 0 & 0 & \dots & 0 \\ 0 & \sigma_u^2 & 0 & \dots & 0 \\ & & & \vdots & \\ 0 & 0 & 0 & \dots & \sigma_u^2 \end{bmatrix}$$

- In less matrix notation:
  - ▶  $\text{var}(u_i) = \sigma_u^2$  for all  $i$  (constant variance)
  - ▶  $\text{cov}(u_i, u_j) = 0$  for all  $i \neq j$  (implied by iid)



# Sampling variance for OLS estimates

- Under assumptions 1-5, the sampling variance of the OLS estimator can be written in matrix form as the following:

$$\text{var}[\widehat{\boldsymbol{\beta}}] = \sigma_u^2 (\mathbf{X}'\mathbf{X})^{-1}$$

- This symmetric matrix looks like this:

	$\widehat{\beta}_0$	$\widehat{\beta}_1$	$\widehat{\beta}_2$	...	$\widehat{\beta}_k$
$\widehat{\beta}_0$	$\text{var}[\widehat{\beta}_0]$	$\text{cov}[\widehat{\beta}_0, \widehat{\beta}_1]$	$\text{cov}[\widehat{\beta}_0, \widehat{\beta}_2]$	...	$\text{cov}[\widehat{\beta}_0, \widehat{\beta}_k]$
$\widehat{\beta}_1$	$\text{cov}[\widehat{\beta}_1, \widehat{\beta}_0]$	$\text{var}[\widehat{\beta}_1]$	$\text{cov}[\widehat{\beta}_1, \widehat{\beta}_2]$	...	$\text{cov}[\widehat{\beta}_1, \widehat{\beta}_k]$
$\widehat{\beta}_2$	$\text{cov}[\widehat{\beta}_2, \widehat{\beta}_0]$	$\text{cov}[\widehat{\beta}_2, \widehat{\beta}_1]$	$\text{var}[\widehat{\beta}_2]$	...	$\text{cov}[\widehat{\beta}_2, \widehat{\beta}_k]$
$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\ddots$	$\vdots$
$\widehat{\beta}_K$	$\text{cov}[\widehat{\beta}_k, \widehat{\beta}_0]$	$\text{cov}[\widehat{\beta}_k, \widehat{\beta}_1]$	$\text{cov}[\widehat{\beta}_k, \widehat{\beta}_2]$	...	$\text{var}[\widehat{\beta}_k]$

# Inference in the general setting

- Under assumption 1-5 in large samples:

$$\frac{\widehat{\beta}_j - \beta_j}{\widehat{SE}[\widehat{\beta}_j]} \sim N(0, 1)$$

- In small samples, under assumptions 1-6,

$$\frac{\widehat{\beta}_j - \beta_j}{\widehat{SE}[\widehat{\beta}_j]} \sim t_{n-(k+1)}$$

- Thus, under the null of  $H_0 : \beta_j = 0$ , we know that

$$\frac{\widehat{\beta}_j}{\widehat{SE}[\widehat{\beta}_j]} \sim t_{n-(k+1)}$$

- Here, the estimated SEs come from:

$$\widehat{\text{var}}[\widehat{\boldsymbol{\beta}}] = \widehat{\sigma}_u^2 (\mathbf{X}'\mathbf{X})^{-1}$$

$$\widehat{\sigma}_u^2 = \frac{\widehat{\mathbf{u}}'\widehat{\mathbf{u}}}{n - (k + 1)}$$

# Covariance matrix in R

- We can access this estimated covariance matrix,  $\widehat{\sigma}_u^2(\mathbf{X}'\mathbf{X})^{-1}$ , in R:

```
vcov(mod)
```

```
##           (Intercept)  exports      age      male
## (Intercept)  0.0004766593  1.164e-07 -7.956e-06 -6.676e-05
## exports      0.0000001164  1.676e-09 -3.659e-10  7.283e-09
## age          -0.0000079562 -3.659e-10  2.231e-07 -7.765e-07
## male        -0.0000667572  7.283e-09 -7.765e-07  1.909e-04
## urban_dum   -0.0000965843 -4.861e-08  7.108e-07 -1.711e-06
## malaria_ecology -0.000069094 -2.124e-08  2.324e-10 -1.017e-07
##           urban_dum malaria_ecology
## (Intercept) -9.658e-05      -6.909e-06
## exports      -4.861e-08      -2.124e-08
## age          7.108e-07      2.324e-10
## male        -1.711e-06      -1.017e-07
## urban_dum    2.061e-04      2.724e-09
## malaria_ecology 2.724e-09      7.590e-07
```

# Standard errors from the covariance matrix

- Note that the diagonal are the variances. So the square root of the diagonal is are the standard errors:

```
sqrt(diag(vcov(mod)))
```

```
##      (Intercept)      exports      age      male
##      0.02183253      0.00004094      0.00047237      0.01381627
##      urban_dum malaria_ecology
##      0.01435491      0.00087123
```

```
coef(summary(mod))[, "Std. Error"]
```

```
##      (Intercept)      exports      age      male
##      0.02183253      0.00004094      0.00047237      0.01381627
##      urban_dum malaria_ecology
##      0.01435491      0.00087123
```

### III. Estimating Equations and Empirical Results

#### A. OLS Estimates

We begin by estimating the relationship between the number of slaves that were taken from an individual's ethnic group and the individual's current level of trust. Our baseline estimating equation is:

$$(1) \quad trust_{i,e,d,c} = \alpha_c + \beta slave\ exports_e + \mathbf{X}'_{i,e,d,c} \Gamma + \mathbf{X}'_{d,c} \Omega + \mathbf{X}'_e \Phi + \varepsilon_{i,e,d,c}$$

# Wrapping up

- You have the full power of matrices.
- Key to writing the OLS estimator and discussing higher level concepts in regression and beyond.
- Next week: diagnosing and fixing problems with the linear model.